

Alpha-Stable (α -Stable) Probability Laws and Generalized Central Limit Theorem

Classical treatise: B.V. Gnedenko and A.N. Kolmogorov, *Limit Distributions for Sums of Independent Random Variables*.

Classical treatise: W. Feller, *An Introduction to Probability Theory and Its Applications*, Volume 2 ..

Modern exposition: G. Samorodnitsky and M. Taquq, *Stable Non-Gaussian Random Processes: Stochastic Models with Infinite Variance*.

Practical guide: J. Nolan, *Stable Distributions*.

- Limit Theorems: How large-scale random phenomena create strict, nonrandom regularity in their collective action.

1. Stable probability laws

1.1. Definition of stable r.v.'s.

X stable if $\exists b_n > 0, a_n$

$$S_n = \sum_{j=1}^n X_j \stackrel{d}{=} a_n + b_n X \quad (1.1)$$

where X_1, X_2, \dots i.i.d., $X_i \stackrel{d}{=} X$. X strictly stable if $a_n = 0$.

Theorem (Feller). $b_n = n^{1/\alpha}$, $0 < \alpha \leq 2$.

Example 1. Gaussian, zero mean.

$$p_X(x) = \frac{1}{\sqrt{4\pi\sigma}} e^{-x^2/4\sigma^2}, \quad \hat{p}_X(k) = e^{-\sigma^2 k^2} \quad (1.2)$$

Note: $\text{Var}\{X\} = 2\sigma^2$. We compare characteristic functions (CF) on the left and right hand sides of Eq.(1.1) **Reminder:** CF as the Fourier transform of the PDF

$$\text{l.h.s.: } \left\langle e^{ik \sum_{j=1}^n X_j} \right\rangle = \left\langle e^{ikX_1 + ikX_2 + \dots} \right\rangle = \left\langle e^{ikX} \right\rangle^n = e^{-n\sigma^2 k^2}$$

$$\begin{aligned} \text{r.h.s.: } \left\langle e^{ik(a_n + b_n X)} \right\rangle &= e^{ika_n} \hat{p}_X(b_n k) = e^{ika_n - \sigma^2 (b_n k)^2} \\ &\Rightarrow a_n = 0, \quad b_n = \sqrt{n} \end{aligned}$$

Strictly stable.

Example 2. Gaussian, non-zero mean.

$$p_X(x) = \frac{1}{\sqrt{4\pi\sigma}} e^{-(x-a)^2/4\sigma^2}, \quad \hat{p}_X(k) = e^{ika - \sigma^2 k^2} \quad (1.3)$$

Stable.

Example 3. Cauchy.

$$p_X(x) = \frac{1}{\pi} \frac{\sigma}{x^2 + \sigma^2} \Rightarrow \hat{p}(k) = e^{-\sigma|k|} \quad (1.4)$$

$$\text{l.h.s. } \left\langle e^{ik \sum_{j=1}^n X_j} \right\rangle = e^{-n\sigma|k|}$$

$$\text{r.h.s. } \left\langle e^{ik(a_n + b_n X)} \right\rangle = e^{ika_n} \hat{p}(b_n k) = e^{ika_n - \sigma b_n |k|}$$

$$\Rightarrow a_n = 0 \quad , \quad b_n = n$$

Stable.

Example 4. Laplace distribution.

$$p_X(x) = \frac{1}{2\sigma} e^{-\frac{|x|}{\sigma}} \Rightarrow \hat{p}(k) = \frac{1}{1 + \sigma^2 k^2} \quad (1.5)$$

$$\text{l.h.s. } \left\langle e^{ik \sum_{j=1}^n X_j} \right\rangle = \frac{1}{(1 + \sigma^2 k^2)^n}$$

$$\text{r.h.s. } \left\langle e^{ik(a_n + b_n X)} \right\rangle = e^{ika_n} \hat{p}(b_n k) = \frac{e^{ika_n}}{1 + \sigma^2 b_n^2 k^2}$$

Non-stable.

1.2. Equivalent definition of strictly stable r.v.'s

X strictly stable if $\forall \sigma_1, \sigma_2 \exists \sigma, \sigma_1^\alpha + \sigma_2^\alpha = \sigma^\alpha$,

$$\sigma_1 X_1 + \sigma_2 X_2 \stackrel{d}{=} \sigma X \quad , \quad (1.6)$$

where $X_1 \stackrel{d}{=} X_2 = X$.

1.3. Solution for strictly stable

$$\hat{p}(\sigma_1 k) \hat{p}(\sigma_2 k) = \hat{p}(\sigma k) \quad . \quad (1.7)$$

Solution:

$$\hat{l}_\alpha(k) = e^{-|k|^\alpha} \quad . \quad (1.8)$$

Remark. The function e^{-k^α} is not possible, since CF must obey $\hat{p}^*(k) = \hat{p}(-k)$.

Question: Which α 's are possible ?

1. Normalization of the PDF. $\hat{l}_\alpha(0) = 1 \Rightarrow \alpha > 0$

2. Suppose $\alpha > 2$. Then $\langle x^2 \rangle = \int_{-\infty}^{\infty} x^2 l_\alpha(x) dx = - \left. \frac{d^2 \hat{l}_\alpha(k)}{dk^2} \right|_{k=0} = 0$.

Conclusion. $0 < \alpha \leq 2$.

Examples. $\alpha = 2$, $\alpha = 1$ (Gauss and Cauchy).

1.4. General solution of Eq.(1.1) (Lévy; Khintchine; Gnedenko, Kolmogorov):

$$p_{\alpha,\beta}(k; \mu, \sigma) = \exp \left[i\mu k - \sigma^\alpha |k|^\alpha (1 - i\beta \text{sign}(k) \omega(|k|, \alpha)) \right], \quad (1.9)$$

where

$$\omega(k, \alpha) = \begin{cases} \tan(\pi\alpha/2) & \text{if } \alpha \neq 1, \quad 0 < \alpha < 2 \\ -(2/\pi) \ln |k| & \text{if } \alpha = 1 \end{cases} .$$

4-parametric: $0 < \alpha \leq 2$ (**Levy index**), $-1 \leq \beta \leq 1$ (**skewness parameter**), $\sigma > 0$ (**scale parameter**), μ real (**shift parameter**).

1.5. Properties of the PDF

1.5.1. **α and β are important ! μ and δ are less important** (γ and δ can be eliminated by proper shift and scale transformation:

$$p_{\alpha,\beta}(x; \mu, \sigma) = \frac{1}{\sigma} p_{\alpha,\beta} \left(\frac{x - \mu}{\sigma}; 0, 1 \right), \quad \alpha \neq 1 \text{ or } \alpha = 1, \beta = 0 \quad . \quad (1.10)$$

Notation: $p_{\alpha,\beta}(x; 0, 1) \equiv p_{\alpha,\beta}(x)$

1.5.2. **Reflection property:**

$$p_{\alpha,-\beta}(x) = p_{\alpha,\beta}(-x) \quad (\text{trivially checked with Eq.(1.9)}) . \quad (1.11)$$

1.5.3. **The sum of two stable variables with the same Levy index α**

$$\mu = \mu_1 + \mu_2 \quad (1.12a)$$

$$\sigma^\alpha = \sigma_1^\alpha + \sigma_2^\alpha \quad (1.12b)$$

(generalization of the rule of summing variances)

$$\beta = \frac{\beta_1 \sigma_1^\alpha + \beta_2 \sigma_2^\alpha}{\sigma^\alpha}, \quad -1 \leq \beta \leq 1 . \quad (1.12c)$$

1.5.4. **Stable laws are continuous and have derivatives of all orders at every point .**

1.5.5. **Unimodality** (single max)

1.6. Symmetric stable laws: $\beta=0$ $p_{\alpha,0}(x) = p_{\alpha,0}(-x)$

1.6.1. CF for symmetric stable law:

$$p_{\alpha,0}(k) \equiv p_{\alpha}(k) = \exp\left(-|k|^{\alpha}\right) .$$

$$p_{\alpha}(k) \approx 1 - |k|^{\alpha} + \frac{1}{2}|k|^{2\alpha} - \dots \quad k \rightarrow 0 .$$

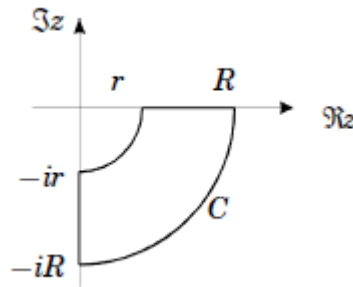
1.6.2. Asymptotics of the PDF

Due to symmetry we may consider $x > 0$, only.

$$l_{\alpha}(x) = \frac{1}{2\pi} \int_{-\infty}^{\infty} e^{ikx - |k|^{\alpha}} dk = \frac{1}{\pi} \operatorname{Re} \int_0^{\infty} e^{-ikx - k^{\alpha}} dk . \quad (1.13)$$

Let us consider the integral in the complex plane,

$$\oint_C e^{-ikz - z^{\alpha}} dz = 0 .$$



Let $r \rightarrow 0$, $R \rightarrow \infty$. Then, according to *Jordan's lemma*, the integral along the large arc vanishes. The integral along the small arc vanishes as well. As the result, for our integral along the real axis we have

$$\int_0^{\infty} e^{-ikx - k^{\alpha}} dk = - \int_{-i\infty}^{-i0} e^{-ikx - k^{\alpha}} dk = \langle ik = \kappa \rangle = - \frac{1}{i} \int_{\infty}^0 e^{-x\kappa - \kappa^{\alpha}} e^{-i\pi\alpha/2} d\kappa =$$

$$-i \int_0^{\infty} e^{-x\kappa - \kappa^{\alpha}} e^{-i\pi\alpha/2} d\kappa . \quad (1.14)$$

Plug (1.14) into (1.13), change variable $x\kappa = y$, and expand exponential function in the integrand. As the result, Eq.(1.13) reads

$$l_{\alpha}(x) = \frac{1}{\pi x} \operatorname{Re} \left\{ -i \int_0^{\infty} e^{-y - \frac{y^{\alpha}}{x^{\alpha}}} e^{-i\pi\alpha/2} dy \right\} = \frac{1}{\pi x} \operatorname{Re} \left\{ -i \int_0^{\infty} e^{-y} \left[1 - \frac{y^{\alpha}}{x^{\alpha}} e^{-i\pi\alpha/2} + O\left(\frac{y^{2\alpha}}{y^{2\alpha}}\right) \right] dy \right\} =$$

Integrate term by term, the first term gives zero

$$\approx \frac{1}{\pi x^{1+\alpha}} \operatorname{Re} \left\{ +ie^{-i\pi\alpha/2} \int_0^{\infty} e^{-y} y^{\alpha} dy \right\} \Rightarrow$$

Finally,

$$p_{\alpha}(x) \approx \frac{C(\alpha)}{|x|^{1+\alpha}}, \quad x \rightarrow \pm\infty \quad (1.15)$$

$$C(\alpha) = \frac{1}{\pi} \sin\left(\frac{\pi\alpha}{2}\right) \Gamma(1+\alpha). \quad (1.16)$$

1.6.3. Asymptotics of the cumulative probability

$$L_{\alpha}(x) = \int_{-\infty}^x l_{\alpha}(y) dy \sim \begin{cases} \frac{C(\alpha)}{\alpha |x|^{\alpha}}, & x \rightarrow -\infty \\ 1 - \frac{C(\alpha)}{\alpha x^{\alpha}}, & x \rightarrow \infty \end{cases}. \quad (1.17)$$

1.6.4. Moments

$$\langle |x|^q \rangle = \int_{-\infty}^{\infty} |x|^q l_{\alpha}(x) dx = \begin{cases} C(q; \alpha), & q < \alpha \\ \infty, & q \geq \alpha \end{cases}. \quad (1.18)$$

Examples. In terms of elementary functions:

Example 1. Gauss

$$l_2(x) = \frac{1}{\sqrt{4\pi}} e^{-x^2/4}.$$

Example 2. Cauchy

$$l_1(x) = \frac{1}{\pi(1+x^2)}.$$

Remark 2. *Asymmetric densities* $0 < \beta < 1$. The decay at $x \rightarrow \infty$ is the same, but $C(\alpha) \rightarrow C(\alpha, \beta)$.

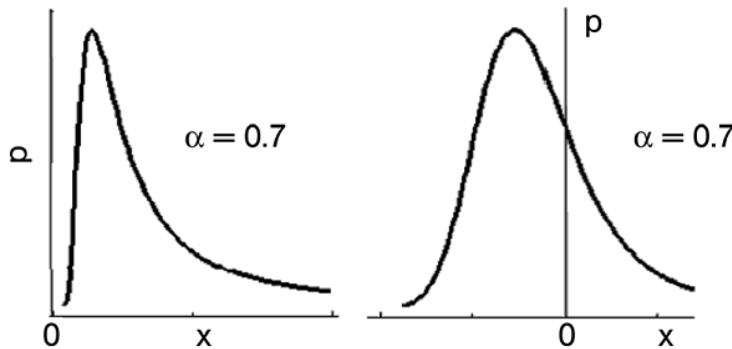
1.7. *Extremal or totally skewed stable laws: $\beta = \pm 1$.*

Go to Eq.(1.9), take $\beta=1$ and put (just for simplicity) $\mu=0$.

$$\hat{l}_{\alpha,1}(k) = \exp \left[-\sigma^{\alpha} |k|^{\alpha} \left(1 - i \operatorname{sign}(k) \tan\left(\frac{\pi\alpha}{2}\right) \right) \right] =$$

$$\begin{aligned}
&= \exp \left[-\frac{\sigma^\alpha |k|^\alpha}{\cos\left(\frac{\pi\alpha}{2}\right)} \left(\cos\left(\frac{\pi\alpha}{2}\right) - i \operatorname{sign}(k) \sin\left(\frac{\pi\alpha}{2}\right) \right) \right] = \\
&= \exp \left[-\frac{\sigma^\alpha |k|^\alpha}{\cos\left(\frac{\pi\alpha}{2}\right)} \left(\cos\left(\frac{\pi\alpha}{2} \operatorname{sign}(k)\right) - i \sin\left(\operatorname{sign}(k) \frac{\pi\alpha}{2}\right) \right) \right] = \\
&= \exp \left[-\frac{\sigma^\alpha |k|^\alpha}{\cos(\pi\alpha/2)} e^{-i \operatorname{sign}(k) \pi\alpha/2} \right]. \tag{1.19}
\end{aligned}$$

Similar to derivation in (1.13-1.15) it can be shown that $l_{\alpha,1}(x) \sim x^{-1-\alpha}$, $x \rightarrow \infty$



Sketch of one-sided (left) and two-sided (right) totally skewed (extremal) α -stable PDFs with $\beta = 1$.

In terms of elementary function: only for $\alpha = 1/2$, $\beta = 1$ (Lévy-Smirnov distribution).

$$p_{1/2,1}(x) = \begin{cases} \frac{1}{\sqrt{2\pi x^3}} e^{-1/2x} & , x \geq 0 \\ 0 & , x < 0 \end{cases} .$$

2. Generalized Central Limit Theorem

Under what conditions on identically distributed independent variables X_1, \dots, X_n, \dots can a limit relation

$$\Pr \left\{ \frac{\sum_{i=1}^n X_i - A_n}{B_n} < z \right\} \rightarrow L(z) \quad , \quad n \rightarrow \infty \tag{2.1}$$

hold, where A_n and B_n are constants, and what kind of limit laws can appear ?

2.1. Reminder: Central Limit Theorem

2.1.1. A pedestrian approach to CLT. Random walk picture.

Random walk is succession of random steps.

What is the distribution of the sum of r.v.'s ?

Example 1. Gaussian. X_1, \dots, X_n i.i.d. Gaussian

$$p_X(x) = \frac{1}{\sqrt{4\pi\sigma^2}} \exp\left[-\frac{x^2}{4\sigma^2}\right] \quad \hat{p}(k) = e^{-\sigma^2 k^2} \quad p_n(x) = ? \quad . \quad (2.2)$$

$$E[X] = 0 \quad , \quad \text{Var}[X] := E\left[(X - E[X])^2\right] := \langle (x - \langle x \rangle)^2 \rangle = 2\sigma^2 \quad .$$

Let us calculate the PDF $p_n(x)$ of the sum $\sum_{j=1}^n X_j$. We have already seen that it can be easily done by making use of the CFs:

$$\begin{aligned} \hat{p}_n(k) &= \left\langle e^{ik \sum_{j=1}^n X_j} \right\rangle = \left\langle e^{ikX_1 + ikX_2 + \dots + ikX_n} \right\rangle = \prod_{j=1}^n \hat{p}_j(k) = \\ &= e^{-n\sigma^2 k^2} \quad \div \quad p_n(x) = \frac{1}{2\pi} \int_{-\infty}^{\infty} e^{-n\sigma^2 k^2} e^{-ikx} dk = \frac{1}{2\pi} \int_{-\infty}^{\infty} e^{-n\sigma^2 \left(k^2 + 2k \frac{ix}{2n\sigma^2} - \frac{x^2}{4n^2\sigma^2} + \frac{x^2}{4n^2\sigma^2} \right)} dk = \\ &= \frac{e^{-\frac{x^2}{4n\sigma^2}}}{2\pi} \int_{-\infty}^{\infty} \exp\left[-n\sigma^2 \left(k - \frac{ix}{2n\sigma^2} \right)^2\right] dk = \frac{e^{-\frac{x^2}{4n\sigma^2}}}{\sqrt{4\pi n\sigma^2}} \Rightarrow \text{also Gauss, but} \end{aligned}$$

$$\boxed{\text{Var}\{X\} \rightarrow \text{Var}\{X\} \times n} \quad (2.3)$$

Example 2. The same Gaussian, but in a different way:

$$\begin{aligned} p_n(x) &= \frac{1}{2\pi} \int_{-\infty}^{\infty} e^{-ikx} \left[e^{-\sigma^2 k^2} \right]^n dk = \langle e^{-\xi} = 1 - \xi + \frac{\xi^2}{2!} - \dots, \quad -\infty < \xi < \infty \rangle = \\ &= \frac{1}{2\pi} \int_{-\infty}^{\infty} e^{-ikx} \left[1 - \sigma^2 k^2 + \dots \right]^n dk = \frac{1}{2\pi} \int_{-\infty}^{\infty} e^{-ikx} e^{n \log(1 - \sigma^2 k^2 + \dots)} dk = \end{aligned}$$

$$= \langle \log(1-\xi) = -\left(\xi + \frac{\xi^2}{2} + \frac{\xi^3}{3} + \dots\right), \quad -1 \leq \xi < 1 \rangle \approx \frac{1}{2\pi} \int_{-\infty}^{\infty} e^{-ikx - \sigma^2 k^2 n} dk \quad (2.4)$$

We got the same result while retaining only the second term in the expansion of the characteristic function.

Conclusion. Gaussian \rightarrow Gaussian, $\boxed{\text{Var}\{X\} \rightarrow \text{Var}\{X\} \times n}$

Example 3. Laplace. X_1, \dots, X_n i.i.d.

$$p_X(x) = \frac{1}{2\sigma} e^{-\frac{|x|}{\sigma}} \quad \div \quad \hat{p}(k) = \frac{1}{1 + \sigma^2 k^2} \quad (2.5)$$

$$E[X] = 0, \quad \text{Var}[X] = E\left[(X - E[X])^2\right] = \langle (x - \langle x \rangle)^2 \rangle = 2\sigma^2.$$

Let us calculate the PDF of the sum $\sum_{j=1}^n X_j$.

$$\begin{aligned} p_n(x) &= \frac{1}{2\pi} \int_{-\infty}^{\infty} e^{-ikx} \left[\frac{1}{1 + \sigma^2 k^2} \right]^n dk \approx \frac{1}{2\pi} \int_{-\infty}^{\infty} e^{-ikx} [1 - \sigma^2 k^2 + \dots]^n dk = \\ &= \frac{1}{2\pi} \int_{-\infty}^{\infty} e^{-ikx} e^{n \ln(1 - \sigma^2 k^2 + \dots)} dk \approx \frac{1}{2\pi} \int_{-\infty}^{\infty} e^{-ikx - n\sigma^2 k^2} dk \\ &= \frac{1}{2\pi} \int_{-\infty}^{\infty} e^{-nb^2 \left(k^2 + 2k \frac{ix}{2nb^2} - \frac{x^2}{4n^2 b^4} + \frac{x^2}{4n^2 b^4} \right)} dk = \\ &= \frac{e^{-\frac{x^2}{4nb^2}}}{2\pi} \int_{-\infty}^{\infty} \exp \left[-nb^2 \left(k + \frac{ix}{2nb^2} \right)^2 \right] dk = \frac{e^{-\frac{x^2}{4nb^2}}}{\sqrt{4\pi nb^2}} \quad (2.6) \end{aligned}$$

Conclusion. Laplace \rightarrow Gaussian, $\boxed{\text{Var}\{X\} \rightarrow \text{Var}\{X\} \times n}$

2.1.2. Central Limit Theorem

X_1, \dots, X_n , $p_X(x)$ with $\text{Var}\{X\} < \infty$

(i) *the summands are i.i.d.*

(ii) *Var{X} finite*

$$\hat{p}(k) \approx 1 - \sigma^2 k^2 + o(k^2), \quad \text{Var}\{X\} = -\left. \frac{d^2 \hat{p}(k)}{dk^2} \right|_{k=0} = 2\sigma^2 < \infty$$

$$p_n(x) \rightarrow \frac{1}{\sqrt{4\pi n\sigma^2}} \exp\left(-\frac{x^2}{4n\sigma^2}\right), \quad n \rightarrow \infty, \quad \boxed{\text{Var}\{X\} \rightarrow \text{Var}\{X\} \times n}$$

The distribution of sum of n independent, identically distributed random variables possessing the mean μ and the variance $\text{Var}\{X\}$ tends for large n to a Gaussian with the mean $\mu \times n$ and the variance $\text{Var}\{X\} \times n$.

2.1.3. Weakening (i) and (ii)

- Lindeberg Theorem

X_1, \dots, X_n, \dots independent variables with $\text{Var}\{X_1\}, \text{Var}\{X_2\}, \dots, \text{Var}\{X_n\}, \dots$

$$p_n(x) \rightarrow \frac{1}{\sqrt{2\pi \left(\sum_{i=1}^n \text{Var}\{X_i\} \right)}} \exp \left(-\frac{x^2}{2 \left(\sum_{i=1}^n \text{Var}\{X_i\} \right)} \right), \quad n \rightarrow \infty$$

$$\text{Var}\{X\} \rightarrow \sum_{i=1}^n \text{Var}\{X_i\} .$$

- Borderline case

X_1, \dots, X_n, \dots i.i.d. with $p_X(x) \propto x^{-3}$

$$p_n(x) \rightarrow \text{Gauss with } \text{Var} \left\{ \sum_{i=1}^n X_i \right\} \sim n \ln n, \quad n \rightarrow \infty$$

Prefactor depends on the details of the distribution of X .

Paul Levy 1930s: Variance = ∞ ?

2.2. Generalized Central Limit Theorem. Simple example with symmetric distribution.

$X_1, X_2, \dots, X_n, \dots$ i.i.d. $p_X(x), \text{Var}\{X\}$ infinite. Limit distribution ?

Intuitively:

$$\hat{p}(k) \approx 1 - C_\alpha |k|^\alpha + o(k^2), \quad \text{Var}\{X\} = -\left. \frac{d^2 \hat{p}(k)}{dk^2} \right|_{k=0} = \infty, \quad 0 < \alpha < 2 .$$

Take

$$p_X(x) \propto \frac{A}{|x|^{1+\alpha}}, \quad x \rightarrow \infty, \quad 0 < \alpha < 2$$

$$\begin{aligned} \hat{p}(k) &= \int_{-\infty}^{\infty} e^{ikx} p_X(x) dx = 2 \int_0^{\infty} \cos(kx) p_X(x) dx = 1 - 2 \int_0^{\infty} (1 - \cos(kx)) p_X(x) dx \approx \\ &\approx 1 - 2A \int_0^{\infty} \frac{1 - \cos(kx)}{x^{1+\alpha}} dx = 1 - 2A |k|^\alpha \int_0^{\infty} \frac{1 - \cos(y)}{y^{1+\alpha}} dy. \end{aligned}$$

$$\text{Separately: } \int_0^{\infty} \frac{1 - \cos(y)}{y^{1+\alpha}} dy = \langle u = 1 - \cos y, dv = dy / y^{1+\alpha} \parallel du = \sin y dy, v = -1/(\alpha y^\alpha) \rangle =$$

$$= \frac{1}{\alpha} \int_0^{\infty} \frac{\sin y}{y^\alpha} dy = \frac{\pi}{2\Gamma(\alpha+1)\sin(\pi\alpha/2)}.$$

Thus,

$$\hat{p}(k) \approx 1 - \tilde{A} |k|^\alpha + \dots, \quad (2.7)$$

where

$$\tilde{A} = \frac{\pi A}{\Gamma(\alpha+1)\sin(\pi\alpha/2)}.$$

Now, in Eq.(2.7) the same trick as before with the Gaussian, namely, retain only these first two terms in the expansion of the CF $\hat{p}(k)$:

$$\hat{p}_n(k) = (\hat{p}(k))^n \approx (1 - \tilde{A} |k|^\alpha)^n = e^{n \ln(1 - \tilde{A} |k|^\alpha)} \approx e^{-\tilde{A} n |k|^\alpha}. \quad (2.8)$$

Remembering the CF of the symmetric α -stable law:

$$\hat{l}_{\alpha,0}(k) = \exp(-\sigma^\alpha |k|^\alpha), \quad 0 < \alpha < 2,$$

GCLT, particular case. We conclude that for i.i.d. with

$$p_X(x) \propto \frac{A}{|x|^{1+\alpha}}, \quad x \rightarrow \infty, \quad 0 < \alpha < 2,$$

the PDF of $\sum_{i=1}^n X_i$

$$p_n(x) \rightarrow l_\alpha(x; \sigma).$$

where

$$\sigma^\alpha = \tilde{A}n = \frac{\pi An}{\Gamma(\alpha+1)\sin(\pi\alpha/2)} = \frac{An}{C(\alpha)}, \quad C(\alpha) = \frac{1}{\pi} \sin\left(\frac{\pi\alpha}{2}\right) \Gamma(1+\alpha).$$

Now, ATTENTION ! Remembering asymptotics of the α -stable PDF,

$$l_\alpha(x; \sigma) \approx C(\alpha) \frac{\sigma^\alpha}{|x|^{1+\alpha}}, \quad x \rightarrow \pm\infty$$

The limit α -stable PDF has the asymptotics

$$l_\alpha(x; \sigma = \frac{An}{C(\alpha)}) \approx \frac{An}{|x|^{1+\alpha}}, \quad x \rightarrow \pm\infty.$$

2.3. GCLT in words

Stable distributions (as the Gaussian law) are the limit ones for distributions of sums of random variables (have domain of attraction) \Rightarrow appear, when evolution of the system or the result of experiment is determined by the sum of a large number of random quantities

Gauss: attracts $f(x)$ with finite variance:

$$f(x) \propto |x|^{-1-\mu}, \quad 0 < \mu < 2$$

$$\langle x^2 \rangle = \int_{-\infty}^{\infty} x^2 f(x) dx < \infty$$

(Lévy) stable: attract $f(x)$ with infinite variance and the same asymptotic behavior

$$\langle x^2 \rangle = \int_{-\infty}^{\infty} x^2 f(x) dx = \infty, \quad f(x) \propto |x|^{-1-\alpha}, \quad 0 < \alpha < 2$$

