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A single predator charging a herd of prey: effects of self volume and predator–prey decision-making

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Abstract

We study the degree of success of a single predator hunting a herd of prey on a two-dimensional square lattice landscape. We explicitly consider the self volume of the prey restraining their dynamics on the lattice. The movement of both predator and prey is chosen to include an intelligent, decision making step based on their respective sighting ranges, the radius in which they can detect the other species (prey cannot recognise each other besides the self volume interaction): after spotting each other the motion of prey and predator turns from a nearest neighbour random walk into directed escape or chase, respectively. We consider a large range of prey densities and sighting ranges and compute the mean first passage time for a predator to catch a prey as well as characterise the effective dynamics of the hunted prey. We find that the prey's sighting range dominates their life expectancy and the predator profits more from a bad eyesight of the prey than from his own good eye sight. We characterise the dynamics in terms of the mean distance between the predator and the nearest prey. It turns out that effectively the dynamics of this distance coordinate can be captured in terms of a simple Ornstein–Uhlenbeck picture. Reducing the many-body problem to a simple two-body problem by imagining predator and nearest prey to be connected by an effective Hookean bond, all features of the model such as prey density and sighting ranges merge into the effective binding constant.

Keywords: first passage process, diffusion, predator–prey model

(Some figures may appear in colour only in the online journal)

1. Introduction

Every animal must eat in order to survive. For certain predator species this necessarily implies to chase and bring down a sufficient amount of prey. With predators always on the lookout for food, prey must constantly be on the alert. While scattering and zigzagging to confuse the predator is a popular method of herd animals to escape an attacking predator [1, 2], if the escape paths are not well co-ordinated individual prey may also block each other. The self volume effect is also relevant in the hunt of killer cells (macrophages, for instance) in biological organisms attacking bacteria, bacteria colonies or biofilms⁴. In this paper we study the influence of self volume effects on a herd of non-communicating prey with the autonomy of taking decisions on the run, as quantified by the typical time to catch a prey.

In the study of the dynamics of predator–prey systems one is generically interested in the likelihood for the survival of the prey as a function of the parameters of the dynamics of both prey and predator. Prototype mathematical models of predator–prey systems are reaction-diffusion models [3–8], in which both species are assumed to move randomly. In one dimension the survival probability of a diffusing prey exposed to a number of diffusing predators decays as a power law in time [9, 10]. In two dimensions the predators catch the prey with probability one, but the mean life time of the prey is infinite. The survival probability of a lamb in the presence of N lions in two-dimensions decays logarithmically slowly as $\mathcal{S}_N(t) \sim (\ln t)^{-N}$ [10]. In contrast, in dimensions three and above the capture is unsuccessful as a consequence of the transience of random walks [11, 12]. Other features considered in predator–prey models include finite life times of the species [13] or the presence of a third party in the form of a repellent obstructing the predator to reach the prey [14]. Moreover, three groups of species hunting each other were modelled [15], owing to the fact that most animal predators are prey of other animals themselves. Finally, effects of safe havens for prey animals may be considered [16].

While such continuum random walk models revealed various interesting results it is clear that the escape and pursuit dynamics is at least partially deterministic, that is, both predator and prey hunt or escape in some sense intelligently. A way to improve the mathematical modelling is to assume that both species can see each other within a certain radius of vision and try to use this as an advantage in the escape and pursuit process [17, 18]. In such a model the motion consists of random walks which turn into directed ballistic transport once predator and prey spot each other. As shown in [17] the probability to escape can be greatly enhanced if the prey can see the predator and has the possibility to run away. During the pursuit the prey’s movement is superdiffusive. In this scenario a total of three predators may be necessary to catch a single prey [17]. Predators may also optimise their search by sharing information [19]. While the assumption of some level of intelligence certainly makes the model more realistic, there is still one aspect that has up to now been ignored. Namely, in reality prey are impenetrable bodies. Thus, in an abundant population of prey (a lion chasing a herd of antelopes, a wolf charging at a flock of sheep, or a killer cell attacking a bacteria, bacteria colony or biofilm) the prey species may obstruct each other while trying to escape. The self volume (non-phantom) constraint greatly influences the single species and collective dynamics of random walkers [20, 21] leading to qualitative differences in the walkers’

⁴ In the following we use the language of predator–prey systems, keeping in mind the relevance of the model for such cellular systems.

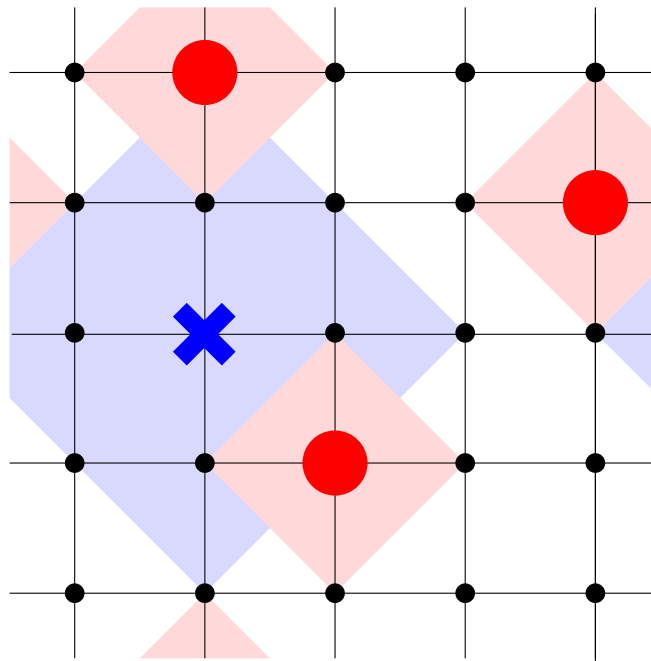


Figure 1. Predator (blue cross) and prey (red dots) on a square lattice. The pale blue and red diamonds represent their respective sighting ranges. Due to their self volume different prey are not allowed to share the same lattice site. Once a prey and the predator meet at the same lattice site the predator is considered to have caught the prey.

motion. Therefore, the dynamics and survival probability in predator prey systems at intermediate and higher prey densities is expected to be equally affected. Recently a herd of prey chased by a pack of predators including self volume effects was studied [22]. As a result the prey's survival time was found to increase if the prey aim for a specific type of clustering.

In this paper we study the success of a single predator hunting a flock of prey on a two-dimensional square lattice with periodic boundary conditions taking into account the prey's self volume. In addition, both species move intelligently in that they can influence their movement by visual perception within their sighting range (figure 1). The paper is structured as follows: First we introduce our model. Next we present the numerical and analytical results for the mean first capture time, which is the time the predator needs to catch the first prey, as a function of prey density and the respective sighting ranges. We find that the mean first capture time as a function of prey density follows a power law. The (non-universal) exponent depends on the sighting ranges of both predator and prey. For the analytical calculations we split the predator's motion into a diffusing part and a ballistic part, representing the search for the prey and the direct chase, respectively. We then present a study of the mean distance between predator and nearest prey, which is found to decrease exponentially in time. Using the mean distance we show that we can capture its dynamics in terms of a simple Ornstein–Uhlenbeck process: the relative motion of predator and nearest prey can thus effectively be viewed to be a random process confined by an harmonic potential. Neglecting all other prey, the model parameters such as sighting ranges and prey density can be absorbed into the associated spring constant.

2. Lattice model

To study the success of a single predator hunting a herd of prey we create an agent-based simulation in which predator and prey move on a two-dimensional square lattice with periodic boundary conditions. Each species has its specific sighting range σ in which it can see the other species as depicted in figure 1. Distances as well as sighting ranges are measured as chemical distances $d = \Delta x + \Delta y$ of the added bond lengths, with lattice spacing a equal to unity. The predator starts from the centre of the lattice and the prey are initially randomly distributed—excluding the centre of the lattice—such that the occupancy of a single site is less or equal to a single prey. Predators and prey move with the autonomy of decision in the following sense. If no prey is in the sighting range of the predator and, for a given prey, the predator is not in its sighting range, both participants perform a nearest neighbour random walk. If a prey comes into the sighting range of the predator, the predator chooses a site randomly, subject to the condition that the distance d to the prey necessarily decreases. Every lattice site that minimises the distance to the prey is chosen with the same probability, lattice sites that increase the distance cannot be chosen. Analogously, if the predator is spotted the prey chooses a site randomly, subject to the condition that the distance to the predator necessarily increases. If two or more prey are within the same distance to the predator the latter chooses randomly which prey to pursue. Due to the self volume of the prey, the prey's motion is restricted. In principle, there exist two possible ways to implement the self volume. Either the prey chooses only from empty sites and always executes a jump as long as there is at least one empty nearest neighbour site. Or the prey blindly chooses a nearest neighbour site but only jumps if the chosen site is unoccupied; otherwise, if it is occupied, the prey retains its location. We chose the latter scenario, as this appears closer to the situation encountered for confused prey or for moving bacteria. Using this update strategy, we simultaneously choose the individual moves for the prey and the predator. In each round of motion updates for the prey we randomly choose a sequence of individuals, thus avoiding any bias among individuals [23]. According to this random sequence we then check whether the individual prey are allowed to jump given the actual positions of all other prey. The motion of the predator takes into account the positions of all prey at the end of the previous update. Once all individual jumps of prey and predator are determined, all positions of the entire predator-prey system are updated simultaneously.

The time unit is chosen arbitrarily and relates to the diffusion constant D

$$\Delta t = \frac{a^2}{4D}, \quad (1)$$

where $a = 1$ is the lattice spacing⁵. After the individual steps of all participants are accomplished, we check if the predator caught a prey. If the first prey is caught the simulation terminates. The mean first capture time and the mean distance are obtained from 10^4 realisations and the first passage density is obtained from 10^6 runs.

3. Mean first capture time

We start by quantifying the success of the predator by computing the mean first capture time $\langle \tau_c \rangle$, that is, the typical time the predator needs to catch the first prey. In mathematical terms this corresponds to the prey's survival time. As one can easily imagine the mean first capture time depends crucially on both sighting ranges σ_{prey} and σ_{pred} as well as on the prey density

⁵ In these units, $D = 1/4$ corresponds to the diffusion coefficient for a single prey or predator moving on the lattice.

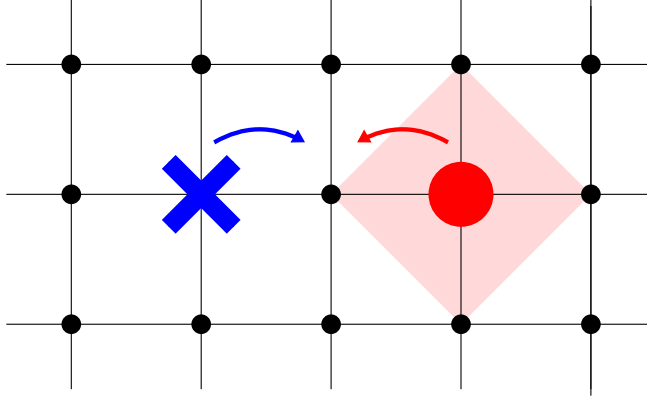


Figure 3. A short-sighted prey ($\sigma_{\text{prey}} = 1$), depicted by the red dot, can get caught by the predator (blue cross) despite his field of vision by random collision, due to simultaneous jumps to the same lattice site.

to catch the first prey. For arbitrary densities, as a result of extensive simulations we find from figure 2 that the mean first capture time as a function of the prey density follows a power law behaviour

$$\langle \tau_c \rangle \sim \varrho^{-\beta(\sigma_{\text{pred}}, \sigma_{\text{prey}})} \quad (2)$$

in which different combinations of sighting ranges lead to different slopes. The exponent β as a function of the sighting range is depicted in figure B2 in the appendix. Furthermore, there appears a crossover between two regimes for larger sighting ranges of the prey, in which we find different slopes for the low and intermediate density range and the high density range; see, for instance, the square symbols in figures 2(b) and (c).

In more detail, while the predator's sighting range only slightly influences the prey's survival, as shown in figure 2(a), the prey can increase their life expectancy significantly by a finite sighting range of at least two, compare figure 2(b), even in the case of a long sighting range of the predator, see figure 2(c). In both figures 2(b) and (c) a significant variation at intermediate σ values is distinct. We note that a short sighting range of the prey ($\sigma_{\text{prey}} = 1$) has no advantage over a vanishing one. An explanation can be found 'microscopically'. There are two possibilities for a prey to get caught. First, a prey gets stuck and, despite his eyesight, cannot evade the encounter with the predator; or, second, predator and prey simultaneously jump on the same lattice site and collide randomly, see figure 3. With sighting range zero or one a prey cannot foresee a random collision, because the distance decreases instantly from two to zero. Thus, the prey needs at least a sighting range of two to prevent such a situation. These random collisions further lead to the fact that the predator is more successful with an even sighting range $\sigma_{\text{pred}} = 2n$, n being an integer number, than with a higher odd one $\sigma_{\text{pred}} = 2n + 1$. Since the random collision is a natural and frequent way to get caught, we decided to eliminate these effects by treating only odd sighting ranges.

We note that we did not include error bars in our figures. A stochastic variable with exponential (Poissonian) probability density function $p(t) = \tau^{-1}e^{-t/\tau}$ has the mean τ and variance $2\tau^2$. The mean first capture time presented in this section is the first moment of the exponentially distributed first passage density obtained in section 4.3. The standard deviation of this Poissonian process $\sigma = \sqrt{\int_0^\infty t^2 p(t) dt - \tau^2}$ is equal to the mean τ , which is indeed

confirmed from our numerical results with a sample size of 10^4 per data point. Repeated simulations produced practically indistinguishable results.

4. Distribution of first capture times

In comparison to an ensemble of non-interacting random walkers self volume effects and the autonomy of decision-making of the participants limit the possibilities of analytical calculations. We succeeded in calculating the distribution of the time for catching the first prey only in the case of blind prey. As the results are nevertheless instructive we discuss this case here in some detail. The autonomy to switch the mode of motion of the predator can be included by dividing the process into two subprocesses. The first one describes the diffusing predator while looking out for a prey. The second subprocess portrays the direct chase of the prey, which can in fact be considered as a ballistic motion in chemical space such that the predator still has the option of choosing sites in different directions.

4.1. Searching the prey

The first subprocess describes the random motion of the predator while looking out for a prey. According to the model during that time the predator performs a nearest neighbour random walk on the lattice. We are interested in the first passage density function of the predator to find the first prey, that is, until the first prey enters the predator's sighting range. For simplification we use a continuous radial coordinate and ignore the fact that the participants move on a lattice. We assume that there exists an effective radius r_{eff} around the predator in which he will not encounter a prey. This radius has a natural lower bound which is the initial distance between predator and nearest prey (at time $t = 0$), calculated in section 5. If the predator hits this effective radius, he spots the prey and will from there on switch his motion to the direct chase calculated in the next subsection.

We consider the predator as a diffusing particle in two-dimensions and calculate his first passage time to escape a sphere with radius $r_{\text{eff}} - \sigma_{\text{pred}}$. For simplification we let the particle diffuse between concentric spheres with an inner reflecting boundary at radius R_- , which will later tend to zero, and an outer absorbing boundary at radius R_+ , representing the point where the predator spots a prey. R_+ is thus the distance between predator and prey minus the sighting range of the predator. The predator starts inside the interval $R_- < r_0 < R_+$. We will later let r_0 tend to R_- to capture the predator's starting position correctly. The diffusing particle can be described by the radial diffusion equation

$$\frac{\partial p(r, t)}{\partial t} = D \frac{1}{r^2} \frac{\partial}{\partial r} \left(r^2 \frac{\partial}{\partial r} \right) p(r, t) \quad (3)$$

for the probability density function $p(r, t)$ to find the predator at radius r at time t . The initial condition we choose as $p(r, t = 0) = \delta(r - r_0)/(2\pi r_0)$, that is, the particle starts at $r = r_0$. We impose the absorbing boundary condition $p(R_+, t) = 0$ at R_+ and the reflecting boundary condition $-\partial p(r, t)/\partial r|_{R_-} = 0$ at $r = R_-$. After Laplace transform

$$\tilde{f}(s) = \mathcal{L}\{f(t)\}(s) = \int_0^\infty f(t) e^{-st} dt \quad (4)$$

and with $x = r\sqrt{s/D}$ the diffusion equation is reduced to the ordinary differential equation

$$\tilde{p}(x, s) - \frac{1}{x} \frac{\partial \tilde{p}(x, s)}{\partial x} - \frac{\partial^2 \tilde{p}(x, s)}{\partial x^2} = \frac{1}{D} \frac{\delta(x - x_0)}{2\pi x_0}. \quad (5)$$

For $x < x_0$ and $x > x_0$ this is the modified Bessel equation of zero order with known solution $\tilde{p}(x, s) = C_1 I_0(x) + C_2 K_0(x)$ for $x \neq x_0$ [24], where $I_0(x)$ and $K_0(x)$ are the modified Bessel functions of first and second kind. We solve this equation by imposing the continuity condition $\tilde{p}_<(x_0, s) = \tilde{p}_>(x_0, s)$ and the jump-discontinuity

$$-\left. \frac{\partial \tilde{p}_>(x, s)}{\partial x} \right|_{x_0} + \left. \frac{\partial \tilde{p}_<(x, s)}{\partial x} \right|_{x_0} = \frac{1}{2\pi D x_0}, \quad (6)$$

where $\tilde{p}_<(x, s)$ is the solution in the range $x < x_0$ and $\tilde{p}_>(x, s)$ is the solution in the range $x > x_0$. With the shorthand notations $C_\nu(a, b) = I_\nu(a)K_\nu(b) - K_\nu(a)I_\nu(b)$ and $D_{\nu,\pm}(a, b) = I_\nu(a)K_{\nu\pm 1}(b) + K_\nu(a)I_{\nu\pm 1}(b)$ [26] the solution yields in the form

$$\tilde{p}(x, s) = \frac{C_0(x, x_+)}{2\pi D x_0 \left(\frac{C_0(x_0, x_+) C_{-1}(x_0, x_-)}{D_{0,-}(x_0, x_-)} - D_{-1,+}(x_0, x_+) \right)}. \quad (7)$$

If the particle starts at the inner boundary $r_0 = R_-$, corresponding to x_- in the reduced coordinates,

$$\lim_{x_0 \rightarrow x_-} \tilde{p}_>(x, s) = -\frac{C_0(x, x_+)}{2\pi D x_0 D_{-1,+}(x_0, x_+)}. \quad (8)$$

we calculate the flux through the outer boundary as

$$-2\pi x_+ D \left. \frac{\partial \tilde{p}_>(x, s)}{\partial x} \right|_{x_+} = (x_- D_{-1,+}(x_-, x_+))^{-1}. \quad (9)$$

When x_- approaches zero, we therefore find that

$$\lim_{x_- \rightarrow 0} \tilde{\rho}_{\text{search}}(s) = (I_0(x_+))^{-1} = \left(I_0 \left(R_+ \sqrt{s/D} \right) \right)^{-1}, \quad (10)$$

where on the right-hand side we restored the original variables. This is but the first passage time density function in Laplace space of the predator to spot a prey. From that time the predator will chase the prey directly, this part being calculated in the next subsection.

4.2. Chasing the prey

The second subprocess, which describes the predator's movement from the moment of spotting the prey until the prey is caught, can be reduced to a one-dimensional problem. Remember that the decision for every step of the predator is constrained by the following rule: the distance to the prey has to necessarily decrease. For the prey, analogously, the goal is to increase the distance. Consequently after a combined predator and prey step the distance between predator and prey can either stay the same or decrease by one lattice spacing if the chosen site of the prey is already occupied and the prey remains at its site. The first capture time can thus be calculated exactly from the number of times a prey remains at its location. A large sighting range of the prey renders the analysis of the chasing process more difficult as all prey try to escape from the predator and will eventually build a cluster that moves away from the predator. Due to the random order of the updates, one cannot say which of the prey remains sitting. Therefore, we

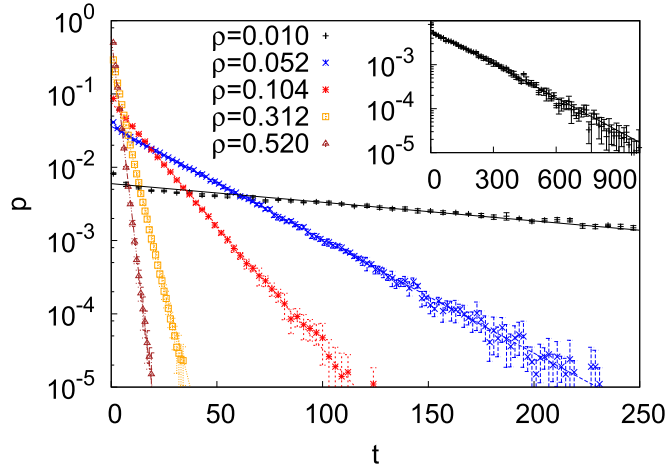


Figure 4. First passage density for the case of blind prey and a short-sighted predator ($\sigma_{\text{pred}} = 1$) for different prey densities. Each data point shows the mean result from 10^6 realisations. The error bars were computed from splitting up the 10^6 independent runs into ten runs of 10^5 runs. Inset: same plot for prey density $\rho = 0.052$ on a larger scale. The lines are exponential fits according to equation (15). The exponent λ as a function of prey density is depicted in figure B1 in appendix B.

confine ourselves to the case of blind prey. In this case the prey undergoes normal diffusion and the predator moves constantly towards the prey. We therefore consider the predator to be a moving cliff towards a diffusing particle, the blind prey. The survival probability of a diffusing particle in presence of a ballistically moving cliff decays exponentially [26],

$$\mathcal{S}(t) \simeq (t/\tau)^{-1/2} e^{-t/\tau}. \quad (11)$$

The associated first passage density $\wp(t) = -d\mathcal{S}(t)/dt$ then becomes

$$\wp_{\text{chase}}(t) \simeq e^{-t/\tau} \left(\frac{(t/\tau)^{-1/2}}{\tau} + \frac{(t/\tau)^{-3/2}}{2\tau} \right). \quad (12)$$

With the Laplace transform $\tilde{\wp}_{\text{chase}}(s) \sim (1 + \tau s)^{-1/2}$ in the long time limit corresponding to a small s expansion, we finally get $\tilde{\wp}_{\text{chase}}(s) \sim (1 + \tau s/2)^{-1}$.

Using the first passage time densities of the subprocesses of search and chase we calculate the total first capture time density function in the next subsection.

4.3. Density of first capture time

The distribution of the first capture time is now given by the convolution of results $\tilde{\wp}_{\text{search}}(s) = (I_0(R + \sqrt{s/D}))^{-1}$ and $\tilde{\wp}_{\text{chase}}(s) \sim (1 + \tau s/2)^{-1}$,

$$\wp(t) = \int_0^t \wp_{\text{search}}(t') \wp_{\text{chase}}(t - t') dt' \quad (13)$$

which designates the probability that the predator spots the first prey at time t' and catches the prey during the time span $t - t'$. In Laplace space this convolution simplifies to the product $\tilde{\wp}(s) = \tilde{\wp}_{\text{search}}(s) \tilde{\wp}_{\text{chase}}(s)$. The inverse Laplace transform can be obtained in the long time limit, corresponding to taking $s \rightarrow 0$. We thus need to invert

$$\tilde{\wp}(s) \simeq (1 + (\kappa + \lambda)s + \kappa\lambda s^2)^{-1}, \quad (14)$$

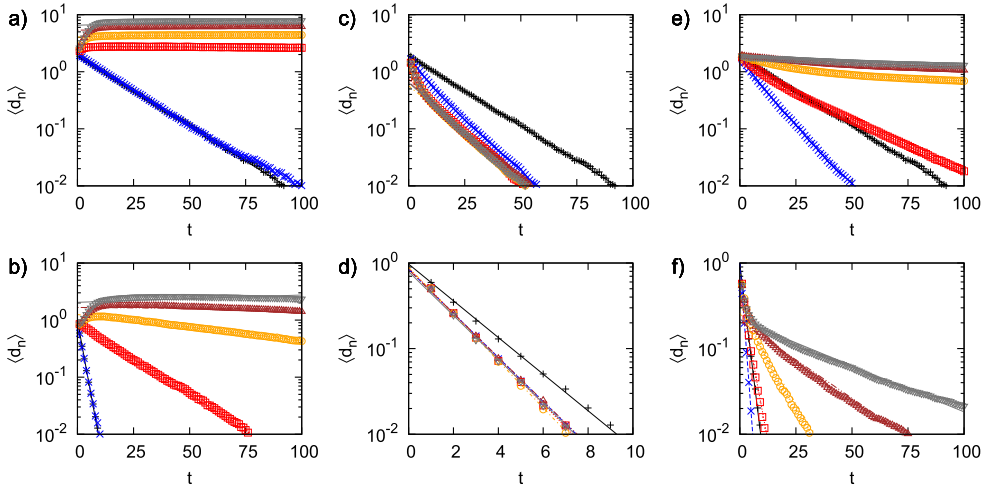


Figure 5. Mean distance between the predator and the nearest prey as function of time, averaged over 10^4 realisations. The upper row (panels (a), (d), (e)) shows the case of a low density $\varrho = 0.104$ and the lower row (panels (b), (d), (f)) represents the case of an intermediate density $\varrho = 0.520$. The two left panels (a) and (b) represent the case of a blind predator, the preys' sighting range decreases from top to bottom ($\sigma_{\text{prey}} = 9, 7, 5, 3, 1, 0$). The two middle panels (c) and (d) represent the case of blind prey. The predator's sighting range increases from top to bottom ($\sigma_{\text{pred}} = 0, 1, 3, 5, 7, 9$). The two panels on the right (e) and (f) show the mean distance in case of identical sighting ranges ($\sigma_{\text{pred}} = \sigma_{\text{prey}} = 0, 1, 3, 5, 7, 9$). They decrease from top to bottom. The lines are exponential fits according to equation (18).

where $\kappa = R_+^2/(4D)$ and $\lambda = \tau/2$. Taking the leading terms for small s , $\varphi(t) \simeq \mathcal{L}^{-1}\{(1 + \Lambda s)^{-1}\}$ with $\Lambda = \kappa + \lambda$, the inverse Laplace transform yields the final result

$$\varphi(t) \simeq \Lambda^{-1}e^{-t/\Lambda}. \quad (15)$$

This density of first capture is thus an exponential distribution, where the rate Λ^{-1} is a function of the prey density and the predator's sighting range. Figure 4 shows the numerical data of our simulation for the case of blind prey and a short sighting range of the predator ($\sigma_{\text{pred}} = 1$). The exponential form (15) agrees quite well with the data over the whole density range.

5. Mean distance between predator and nearest prey

We now turn to study the dynamics of the mean distance between the predator and the nearest prey in more detail. In figure 5 our simulation results for this mean distance are plotted for a low prey density $\varrho = 0.104$ in the upper row (panels (a), (c) and (e)) and for an intermediate density $\varrho = 0.520$ in the lower row (panels (b), (d) and (f)). The distance decreases exponentially in time except for the case when a blind predator is combined with a low prey density (figure 5 (a)) or with a very good eye-sight of the prey (figure 5(b)). In these cases the distance is approximately constant in the shown time window. In case of identical sighting ranges the distance between short-sighted species decreases faster than the distance between blind species. This phenomenon is due to the random collisions explained in section 3.

As intuitively expected, the distance between the predator and the nearest prey decreases faster in the case of a large sighting range of the predator. However when the prey's sighting range is large it softens the decay of the distance. A high prey density also leads to a faster

decay of the distance between the predator and the nearest prey, because it implies more prey-prey obstruction events for the chased prey, and with every one such event the distance is reduced by one lattice spacing.

A naive model that captures the effective interaction between two diffusive particles such as the predator and the nearest prey turns out to be the Ornstein–Uhlenbeck process [25]. It is defined in terms of the stochastic differential equation

$$dx(t) = (e - cx(t)) dt + b dW(t) \quad (16)$$

with non-negative parameters e , b , and c . $W(t)$ denotes the Wiener process [27]. The Ornstein–Uhlenbeck process describes the relaxation of the variable x with initial value $x(t=0) = x_0$ to the mean value e/c in the presence of Gaussian white noise. The first moment is given by the exponential decay

$$\langle x(t) \rangle = \frac{e}{c} [1 - \exp(-ct)] + x_0 \exp(-ct). \quad (17)$$

Comparing the first moment to the observed simulated decay of the mean distance between the predator and the nearest prey (figure 5),

$$\langle d_n \rangle \simeq e^{-\theta t}, \quad (18)$$

we see that the mean distance decreases as a special case of the Ornstein–Uhlenbeck process with vanishing excentricity parameter, $e = 0$.

A popular application of the Ornstein–Uhlenbeck process in physics is a Hookean spring with spring constant k , whose dynamics is highly overdamped with friction coefficient γ in the presence of thermal fluctuations. Therefore we can imagine the predator and the nearest prey to be connected by a Hookean spring and being driven by an external Wiener noise. The corresponding mean relaxes to zero. The equilibrium length of the spring is therefore zero. The bottom of the corresponding harmonic potential thus represents the capture of the prey by the predator. Due to the analogy, the respective sighting ranges and the prey density affect the stiffness of the spring. The spring constant is easily related to the decay rate θ of the mean distance, $\theta(\varrho, \sigma) = k(\varrho, \sigma)/2$. As shown in figure 6 the fitted values for the spring constant display the power law behaviour

$$k \sim \varrho^\nu. \quad (19)$$

The exponent ν as a function of the sighting range is depicted in figure B2 in the appendix. The spring constant corresponds to the slopes of the functions in figure 5, extracted from the exponential fit and plotted as a function of the prey density. It relates to the mean first capture time discussed in section 3 in the following way. The exponential decay of the mean distance between predator and nearest prey has a mean life time related to the decay rate

$$\tau = \frac{1}{\theta}. \quad (20)$$

Since the nearest prey is the one that will get caught, its mean first capture time is related to the mean life time of the mean distance and consequently to the inverse of the decay rate, compare figures 2 and 6.

5.1. Initial distance analysis

We finally mention an analytical approximation for the distance between the predator and the nearest prey. Since we want to capture the whole dynamics we first need to determine the initial distance between predator and nearest prey at time $t = 0$. In the simulation we place the predator in the centre and place the prey randomly around him including the self volume

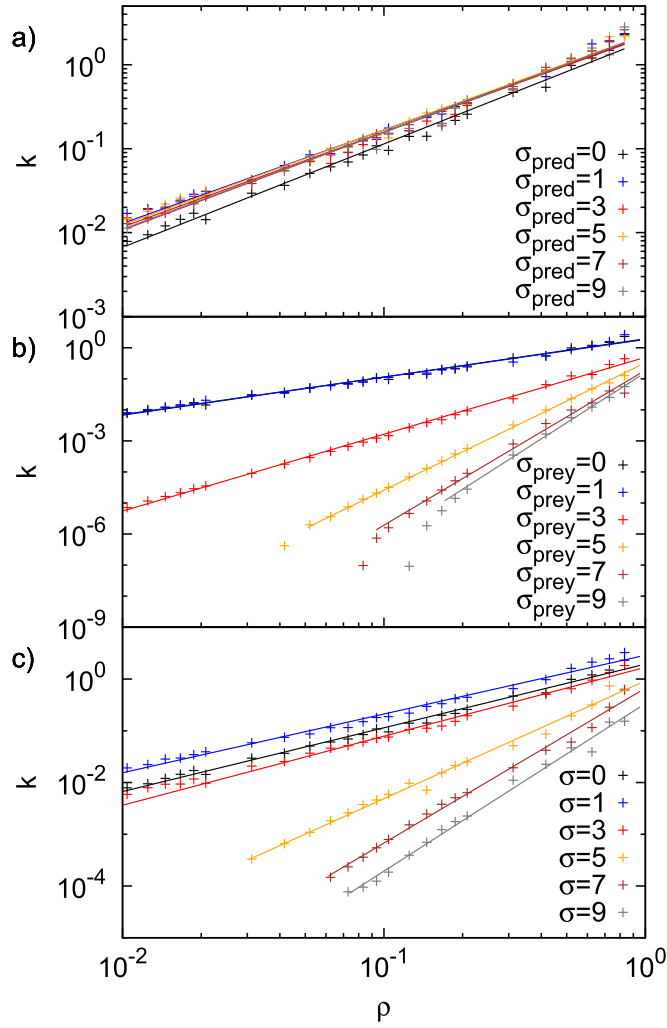


Figure 6. Effective spring constant k of our Ornstein–Uhlenbeck model as function of the prey density for (a) blind prey: $\sigma_{\text{pred}} = 0, 1, 3, 5, 7, 9$, increases from bottom to top. (b) Blind predator: $\sigma_{\text{prey}} = 0, 1, 3, 5, 7, 9$, increases from top to bottom. (c) Identical sighting ranges: $\sigma_{\text{prey}} = \sigma_{\text{pred}} = 0, 1, 3, 5, 7, 9$, increases from top to bottom. The lines are power law fits according to equation (19). The exponent ν as a function of sighting ranges is depicted in figure B2 in appendix B.

interaction. Then we measure the distance between the predator and the nearest prey. In section 4.1 we used an effective radius r_{eff} within which the predator does not encounter a prey. Although we cannot calculate this effective radius a natural lower bound is the initial mean distance $\langle d_n \rangle_{t=0}$ between the predator and the nearest prey. Within this distance there is no prey present and therefore it is impossible for the predator to encounter a prey.

We determine the initial distance between the predator and the nearest prey on a square lattice with edge length L . The predator sits in the centre of the lattice and the prey are randomly distributed on the remaining $N_S = L^2 - 1$ sites. As the prey have a self volume, a lattice site can only be occupied by a single prey. The probability for the distance between

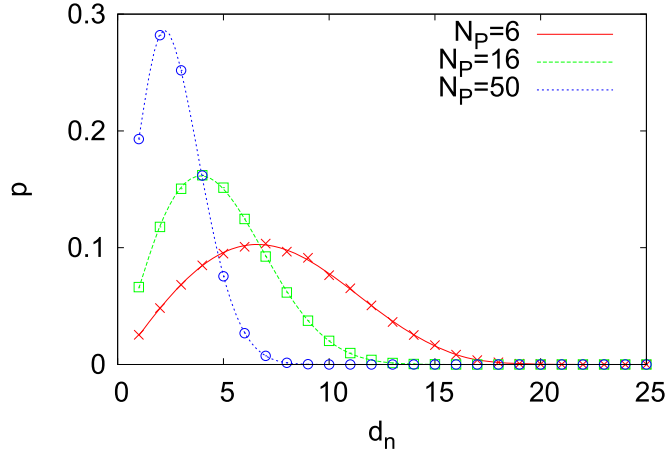


Figure 7. Probability distribution of the initial distance between the predator and the nearest prey for the case of different numbers N_P of prey on a square lattice with edge length $L = 31$. The crosses represent the numerical data, averaged over 10^4 realisations. The lines show the analytical result (22).

predator and nearest prey d_n to be equal to d is

$$P(d_n = d) = P(d_n \geq d) - P(d_n \geq d + 1). \quad (21)$$

We then calculate the probability $P(d_n \geq d)$ using combinatorics. The detailed calculation can be found in appendix A. For the probability function of the distance between predator and nearest prey we obtain

$$P(d_n) = \left[\binom{N_R(d)}{N_P} - \binom{N_R(d+1)}{N_P} \right] / \sum_{i=1}^{d_{\max}} \binom{N_R(d_i)}{N_P}, \quad (22)$$

where we define d_{\max} as the maximal possible distance between the predator and the nearest prey. The expectation value of the initial distance from the predator to the nearest prey, $\langle d_n \rangle = \sum_{d_i=1}^{d_{\max}} p(d_{\min,i}) d_{\min,i}$ then yields in the form

$$\langle d_n \rangle = \frac{\sum_{d_i=1}^{d_{\max}} d_i \left(\binom{N_R(d_i)}{N_P} - \binom{N_R(d_i+1)}{N_P} \right)}{\sum_{d_i=1}^{d_{\max}} \left(\binom{N_R(d_i)}{N_P} - \binom{N_R(d_i+1)}{N_P} \right)}. \quad (23)$$

The probability distribution of the initial distance to the nearest prey is shown in figure 7 and the related initial mean distance as a function of prey density can be seen in figure 8. We simulated both the initial distance distribution and the initial mean distance between the predator and the nearest prey by placing all participants on the lattice under the model conditions with 10^4 iterations. Both analytical and numerical results show excellent agreement in figures 7 and 8.

6. Discussion

We studied the predator–prey dynamics of a single predator hunting a herd of prey on a square lattice with decision-making species. While many predator–prey models deal with collective predation [28–32] or the search for the optimal number of predators given the

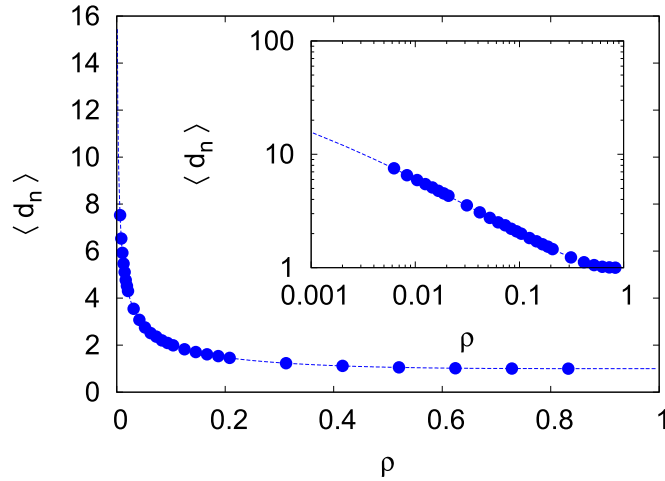


Figure 8. Initial value of the mean distance between the predator and the nearest prey as a function of prey density on a square lattice with edge length $L = 31$. The symbols represent the numerical data, averaged over 10^4 realisations, and the dashed line the analytical result (23).

number of prey [33], we chose a model consisting of one predator and many prey, which is often found in Nature. Solitary hunters such as tigers, bears, or sea turtles often have herd animals as their target. A tiger, for example, hunts a herd of antelopes or a flock of sheep, a bear fishing a salmon out of a stream. Similarly individual killer cells in biological organisms may attack a bacteria, colony of bacteria or a biofilm.

A major ingredient of our model is the self volume of the prey, such that no two prey are allowed on a given lattice site. We showed that in the case of impenetrable prey the predator hunts more successfully if the prey have worse eyesight. Moreover, we found that the predator benefits more from a deterioration of the prey's eyesight than from an improvement of his own eyesight.

While trapping reaction models obtain a minor influence of the prey's long time survival probability by their diffusion constant [5, 6] we found the prey's sighting range and thereby motion predominating their survival probability. Due to self volume interactions the prey are forced to improve their eye-sight, and with a good field of vision can drastically increase their chances of survival even in the range of high densities.

The prey only profit from a sighting range of at least two. A very short eyesight does not at all improve its survival probability with respect to being blind. This is attributed to random collisions between predator and prey. Using a simplified analytic approach we showed that in the long time limit the first passage density of the predator to catch a blind prey decays exponentially in time with a nonlinear dependence of the decay rate on the prey density.

The effective motion during the chase (described in terms of the distance between the predator and the chased prey) can be effectively described as a linear relaxation process in an harmonic potential with a stochastic driving where the density and sighting ranges determine the stiffness of the corresponding Hookean spring. All nonlinear effects entering the motion due to self volume interactions can thus effectively be described with a single parameter.

There exist a range of further open questions. To imitate natural environment one could extend the dynamics by introducing (time or sighting range dependent) waiting times. One could choose different rates of motion for predator and prey as well or even distribute the rates within the prey to simulate old, sick or infant animals. Additionally, many prey live in herds, so one could let the prey be clustered as the initial condition. Last but not least,

communication between the prey is a reasonable thing to assume. Once one of the prey spots the predator, immediately all of them are informed (similar to stamping of rabbits or the cheeping calls of groundhogs), that is, a collective response of prey.

We finally note that random search processes with non-Brownian search dynamics are also widely discussed in literature. While Brownian motion is an advantageous process to find nearby targets [34], it is known that pure stochastic motion leads to oversampling of the area on longer time scales. Hence, the optimal number of encounters with prey can be found by switching between search modes [35, 36]. Representative for such a process is for example the intermittent search strategy which combines phases of slow motion, allowing the searcher to detect the target, and phases of fast motion during which targets cannot be detected [37, 38]. Another widely applicable process concerning optimal search strategies are Lévy flights, which are based on random walk processes with long-tailed jump length distributions and are known to be an efficient strategy for finding a target of unknown place [39, 40]. A species which is known to move in Lévy patterns are wandering albatrosses [41, 42] or marine predators as sharks, bony fishes, sea turtles and penguins [43, 44]. It would thus be interesting to study effects of self volume in these models as well.

Acknowledgments

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Appendix A. Initial distance between predator and nearest prey

We determine the initial distance between the predator and the nearest prey on a two-dimensional square lattice with edge length L . The predator sits in the centre of the lattice and the prey are randomly distributed on the remaining $N_S = L^2 - 1$ sites. As the prey have a self volume, a lattice site can only be occupied by a single prey. The probability for the minimal distance between predator and nearest prey d_n to be equal to d is

$$P(d_n = d) = P(d_n \geq d) - P(d_n \geq d + 1). \quad (\text{A.1})$$

We calculate the probability function $p(d_n \geq d)$ using combinatorics. If $d_n \geq d$ all sites within distance d (up to distance $d - 1$) must be unoccupied. To obtain the number of these sites we count all sites at exactly distance d and add them from distance 1 up to $d - 1$. The number of sites at distance d can be shown to be

$$N(d) = \begin{cases} 4d, & d \leq (L - 1)/2 \\ 4(L - d), & d > (L - 1)/2 \end{cases} \quad (\text{A.2})$$

Counting all empty sites within the distance d from the predator leads to

$$M(d) = \sum_{i=1}^{d-1} N(i). \quad (\text{A.3})$$

That is explicitly

$$M(d) = \begin{cases} \sum_{i=1}^{d-1} 4i = 2d(d - 1), & d \leq (L + 1)/2 \\ N_S - \sum_{i=1}^{L-d} 4i = (L^2 - 1) - 2(L - d)(L - d + 1), & d > (L + 1)/2 \end{cases}. \quad (\text{A.4})$$

Due to the predator sitting in the centre there are in general $N_S = L^2 - 1$ possible sites for the prey to be placed on. Under the assumption that the minimal distance is d , i.e., $M(d)$ sites are vacant, there are $N_R(d) = N_S - M(d)$ remaining sites for the prey. The probability for the minimal distance to be greater or equal d is the number of possibilities to place the prey at the remaining sites $N_R(d)$ over the possibilities to place the prey at sites greater equal every possible distance (1 to d_{\max})

$$P(d_n \geq d) = \binom{N_R(d)}{N_P} / \sum_{i=1}^{d_{\max}} \binom{N_R(d_i)}{N_P}. \quad (\text{A.5})$$

For the probability function of d_n using equation (21) we obtain

$$P(d_n) = \left(\binom{N_R(d)}{N_P} - \binom{N_R(d+1)}{N_P} \right) / \sum_{i=1}^{d_{\max}} \binom{N_R(d_i)}{N_P}, \quad (\text{A.6})$$

where we define d_{\max} as the maximal possible distance between the predator and the nearest prey. It is determined by the number of prey (due to the self volume of the prey) and can be calculated by allocating all prey as greatest distance as possible starting at $d = L - 1$. Then the first fully unoccupied diamond at distance d is the maximal possible distance d_{\max} . There exist the following condition to place all prey $N_P \leq N_S - M(d_{\max})$,

$$\begin{cases} N_P \leq \sum_{i=1}^{L-1-d_{\max}} 4i, & N_P < \frac{N_S}{2} \\ N_P - \frac{L^2-1}{2} \leq \sum_{i=d_{\max}}^{\frac{L-1}{2}} 4i, & N_P \geq \frac{N_S}{2} \end{cases}. \quad (\text{A.7})$$

We then get the maximal possible distance as a function of prey

$$d_{\max} = \begin{cases} \left\lfloor \frac{2L-1-\sqrt{1+2N_P}}{2} \right\rfloor, & N_P < N_S/2 \\ \left\lfloor \frac{\sqrt{1+2(N_S-N_P)}-1}{2} \right\rfloor, & N_P \geq N_S/2 \end{cases}, \quad (\text{A.8})$$

where $\lfloor x \rfloor := \max \{m \in \mathbb{Z} | m \leq x\}$ is the floor function. We now obtain the expectation value of the initial distance from the predator to the nearest prey

$$\langle d_n \rangle = \sum_{d_i=1}^{d_{\max}} P(d_{\min,i}) d_{\min,i} \quad (\text{A.9})$$

such that

$$\langle d_n \rangle = \frac{\sum_{d_i=1}^{d_{\max}} d_i \left(\binom{N_R(d_i)}{N_P} - \binom{N_R(d_i+1)}{N_P} \right)}{\sum_{d_i=1}^{d_{\max}} \left(\binom{N_R(d_i)}{N_P} - \binom{N_R(d_i+1)}{N_P} \right)}. \quad (\text{A.10})$$

Appendix B. Exponents of figures 2, 4 and 6

We here present plots depicting the dependence of the parameter Λ from figure 4 versus the prey density (figure B1) as well as of the scaling exponents β and ν from figures 2 and 6.

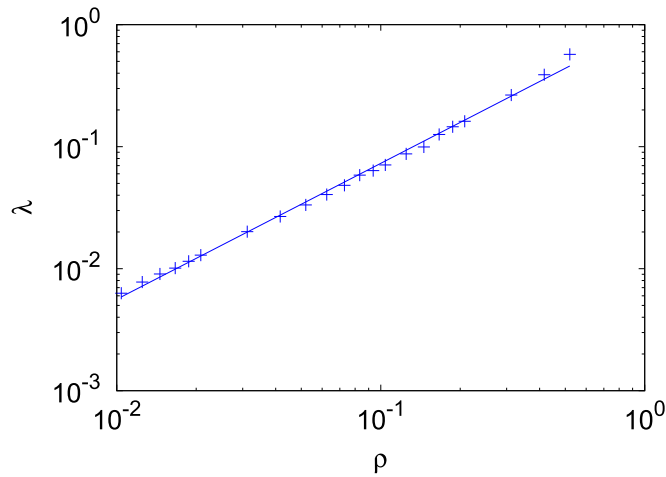


Figure B1. Parameter Λ from the exponential fits in figure 4 as function of prey density.

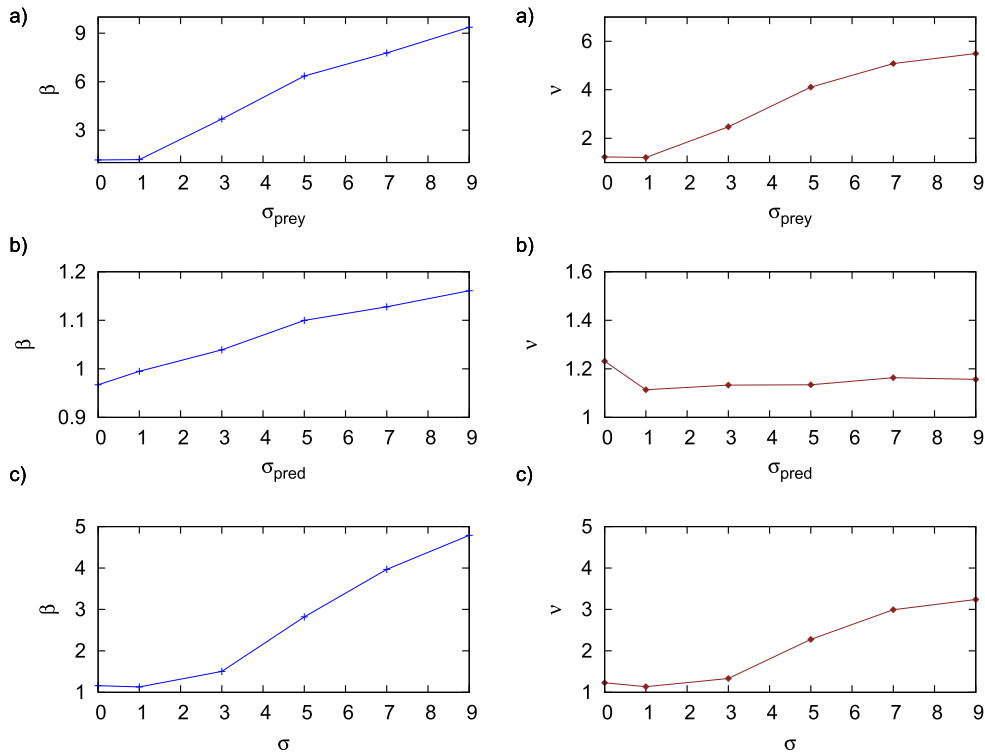


Figure B2. Left: exponents of the power-law fits in figure 2(a) as function of the preys' sighting range in case of a blind predator, (b) as function of the predator's sighting range in case of blind prey, (c) as function of the sighting range in case of identical sighting ranges. Right: exponents of the power-law fits in figure 6(a) as function of the preys' sighting range in case of a blind predator, (b) as function of the predator's sighting range in case of blind prey, (c) as function of the sighting range in case of identical sighting ranges. The lines are meant to guide the eye.

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